## Machine Learning in Bioinformatics <br> Model Selection and Regularization

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MOdeL SELECTION PROBLEM


Linear model class

## MODEL SELECTION



Quadratic model class

## MODEL SELECTION



Polynomial model class

## BIAS-VARIANCE DECOMPOSITION AND TRADEOFF

## BIAS-VARIANCE DECOMPOSITION

- Let $\mathbf{Y}, \mathbf{X}$ and $\epsilon$ be random variables such that $\mathbf{Y}=f(\mathbf{X})+\epsilon$, with $\mathbb{E}[\epsilon]=0$ and $\operatorname{var}[\epsilon]=\sigma^{2}$
- Assume that $\hat{f}_{D}$ has been estimated on some training data $D=(X, y)$, where $X$ is a matrix of $n$ observations from $X$ and $y$ a vector of $n$ observations from $\mathbf{Y}$

■ At a query point $x$ we have

$$
\mathbb{E}_{\mathbf{Y}, D}\left[\left(\mathbf{Y}-\hat{f}_{D}(x)\right)^{2}\right]=\underbrace{\left[\mathbb{E}_{D} \hat{f}_{D}(x)-f(x)\right]^{2}}_{\text {Bias }^{2}}+\underbrace{\mathbb{E}_{D}\left[\hat{f}_{D}(x)-\mathbb{E}_{D} \hat{f}_{D}(x)\right]^{2}}_{\text {Variance }}+\sigma^{2}
$$

■ bias: Is there a bias towards a particular kind of solution (e.g. linear model)? (inductive bias)

■ variance: How much does the estimated model change if you train on a different data set? (overfitting)

## BIAS-VARIANCE DECOMPOSITION

Low variance


High variance


## BIAS-VARIANCE DECOMPOSITION


${ }^{\circ}$ Note that here we average over multiple data sets. On a single data set we might observe bumps when increasing model complexity

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## BIAS-VARIANCE DECOMPOSITION - LESSIONS LEARNED

■ Every model comes with a bias
■ More complex models have a smaller bias but larger variance
■ A bias is required to reduce the variance, but introducing a good bias requires domain knowledge

- Classical statistics often uses unbiased estimators, which is nowadays often questioned

■ Keep in mind: There is no free lunch! ${ }^{1}$

[^0]
## COMPLEXITY MEASURES

## COMPLEXITY OF CLASSIFIERS - VC DIMENSION

## VC-Dimension (Vapnik Chervonenkis)

Let $\mathbb{F}_{p}$ be a set of classifiers on an $n$-dimensional input space. The VC -dimension $\operatorname{VC}\left(\mathbb{F}_{p}\right)$ is defined as the maximum number of points that can be correctly classified by at least one member of $\mathbb{F}_{p}$.

■ Examples:

- Linear classifier on $\mathbb{R}^{p}: \mathrm{VC}=p+1$
- SVM with RBF kernel: $\mathrm{VC}=\infty$
- Neural network with $n_{e}$ edges, $n_{v}$ nodes and sigmoid activation function: $\Omega\left(n_{e}^{2}\right)<\mathrm{VC}<\mathcal{O}\left(n_{e}^{2} n_{v}^{2}\right)$ [Shalev-Shwartz and Ben-David, 2014, Section 20.4]


## COMPLEXITY OF CLASSIFIERS - VC DIMENSION



## MEASURES OF MODEL COMPLEXITY - DF

## Degrees of Freedom (DF) [Efron, 1986]

The degrees of freedom of an estimate $\hat{y}=\hat{f}(X)$ is defined as

$$
\operatorname{df}(\hat{y})=\frac{1}{\sigma^{2}} \sum_{i=1}^{n} \operatorname{cov}\left(\hat{y}_{i}, y_{i}\right)=\frac{1}{\sigma^{2}} \operatorname{tr} \operatorname{cov}(\hat{y}, y),
$$

where

- $X$ denotes a fixed set of $n$ covariates of dimension $p$
- $y=\left(y_{1}, \ldots, y_{n}\right)$ is a vector of $n$ observations from

$$
\mathbf{Y}=f(X)+\epsilon
$$

for some function $f$, assuming $\mathbb{E}[\epsilon]=0$ and $\operatorname{var}[\epsilon]=\sigma^{2}$
${ }^{1} \mathrm{df}$ is normalized by the magnitude of the aleatory uncertainty $\left(\sigma^{2}\right)$

## MEASURES OF MODEL COMPLEXITY - DF

■ Degrees of freedom for the OLS estimate:

$$
\begin{aligned}
\operatorname{df}(\hat{y}) & =\frac{1}{\sigma^{2}} \operatorname{tr} \operatorname{cov}(\hat{y}, y) \\
& =\frac{1}{\sigma^{2}} \operatorname{tr} \operatorname{cov}\left(X\left(X^{\top} X\right)^{-1} X^{\top} y, y\right) \\
& =\frac{1}{\sigma^{2}} \operatorname{tr}\left(X\left(X^{\top} X\right)^{-1} X^{\top}\right) \operatorname{cov}(y, y) \\
& =\operatorname{tr}\left(X\left(X^{\top} X\right)^{-1} X^{\top}\right) \\
& =p
\end{aligned}
$$

■ $\operatorname{df}(\hat{y})=p$, i.e. the number of parameters, assuming independent feature vectors (i.e. columns of $X$ )

■ This result holds for $p<n$

## MEASURES OF MODEL COMPLEXITY - DF

- Ridge regression is defined as

$$
\hat{\theta}=\underset{\theta}{\arg \min }\|y-X \theta\|_{2}^{2}+\lambda\|\theta\|_{2}^{2}
$$

for some regularization strength $\lambda \geq 0$
■ The ridge estimator has

$$
\operatorname{df}(\hat{y})=\sum_{j=1}^{p} \frac{d_{j}^{2}}{d_{j}^{2}+\lambda}
$$

degrees of freedom, where $\left(d_{j}\right)_{j}$ are the singular values of $X$
■ Increasing $\lambda$ decreases model complexity

## MEASURES OF MODEL COMPLEXITY - DF

- There is some criticism about used DF as measure of model complexity [Janson et al., 2015]
- In some cases, we also need $X$ to be random
[Luan et al., 2021]
- We will see other measures when turning to model selection

MODEL SELECTION

## MODEL SELECTION APPROACHES

■ A measure of accuracy or fit, such as the mean squared error (MSE), is not enough: Increasing model complexity will always lead to a better fit

- Estimating a model requires to minimize both
- in-sample-error (loss on training data), and
- out-of-sample-error (generalization error)

■ Cross-validation (CV) estimates generalization error on left-out samples ${ }^{2}$

- Traditional statistics: Combine measure of accuracy (in-sample-error) with a penalty for complexity
${ }^{2}$ Heavy hyperparameter tuning using CV can lead to overfitting and requires to select a final holdout set


## Model selection approaches - LOO-CV

■ Leave-one-out Cross-Validation (LOO-CV) at iteration $i=1,2, \ldots, n$ :

- Compute estimate on data set without the $i$-th sample
- Compute prediction error on the $i$-th sample

■ Report the average prediction error over all $n$ samples
■ PRESS statistic (predicted residual error sum of squares):

$$
\operatorname{PRESS}=\sum_{i=1}^{n}\left(y_{i}-\hat{y}_{-i}\right)^{2}
$$

where $\hat{y}_{-i}$ is the prediction for the $i$-th sample where the model has been estimated on all but the $i$-th sample

## Model Selection approaches - PRESS

■ LOO-CV is very costly for large data sets and complex models
■ $k$-fold CV with $k=5$ or $k=10$ is often used in practice
■ For (ridge) linear regression with mean squared error we can efficiently compute LOO-CV [Cook, 1977]

$$
\begin{aligned}
\operatorname{PRESS} & =\sum_{i=1}^{n}\left(y_{i}-\hat{y}_{-i}\right)^{2} \\
& =\sum_{i=1}^{n} \frac{\left(y_{i}-\hat{y}_{i}\right)^{2}}{\left(1-H_{i i}\right)^{2}}
\end{aligned}
$$

- The matrix

$$
H=X\left(X^{\top} X+\lambda I\right)^{-1} X^{\top}
$$

is called the hat matrix, because it puts a hat on $y$, i.e. $\hat{y}=H y$

## MODEL SELECTION APPROACHES

- LOO-CV is computationally very expensive
- $k$-fold CV is cheaper, but uses a large fraction of the data for testing
- Model performance could be better if this data was used for training
- Overfitting if we use CV for testing too many models (requires final hold out data)
- Can we do model selection by using all data for training?


## Model selection approaches - DF

- Assume again the following model

$$
\mathbf{Y}=f(X)+\epsilon
$$

where $X \in \mathbb{R}^{n \times p}$ is a fixed set of $n$ predictors and $\mathbf{Y} \in \mathbb{R}^{n}$
■ Setup is very similar to the bias-variance decomposition, but $X$ is now fixed

■ Let $\mathbf{Y}_{t} \in \mathbb{R}^{n}$ a vector of $n$ independent observations and $\hat{f}_{\mathbf{Y}_{t}}$ an estimate on the training set $\left(X, \mathbf{Y}_{t}\right)$, then [Efron, 1986]

$$
\underbrace{\mathbb{E}_{\mathbf{Y}, \mathbf{Y}_{t}}\left\|\mathbf{Y}-\hat{f}_{\mathbf{Y}_{t}}(X)\right\|_{2}^{2}}_{\text {expected prediction error }}=\underbrace{\mathbb{E}_{\mathbf{Y}_{t}}\left\|\mathbf{Y}_{t}-\hat{f}_{\mathbf{Y}_{t}}(X)\right\|_{2}^{2}}_{\text {expected training error }}+2 \sigma^{2} \mathrm{df}(\hat{f})
$$

## MODEL SELECTION APPROACHES - DF

■ This motivates the following model selection criterium [Mallows, 2000]

$$
\underbrace{\left\|y_{t}-\hat{f}_{y_{t}}(X)\right\|_{2}^{2}}_{\text {training error }}+\underbrace{2 \sigma^{2} \mathrm{df}(\hat{f})}_{\text {complexity penalty }}
$$

■ The more complex a model, the larger the penalty
■ If two models fit the data equally well, we select the simpler one (Occam's razor)

## Model selection approaches - Bayes Approach

■ Assume we have a set of models $\left(m_{i}\right)_{i}$
■ In a probabilistic setting we evaluate the probability of a model $m_{i}$ given data $x$, i.e. using Bayes theorem

$$
\operatorname{pr}\left(m_{i} \mid x\right)=\frac{\operatorname{pr}\left(x \mid m_{i}\right) \operatorname{pr}\left(m_{i}\right)}{\sum_{j} \operatorname{pr}\left(x \mid m_{j}\right) \operatorname{pr}\left(m_{j}\right)}=\frac{\operatorname{pr}\left(x \mid m_{i}\right) \operatorname{pr}\left(m_{i}\right)}{\operatorname{pr}(x)}
$$

- We compare two models $m_{i}$ and $m_{j}$ using

$$
\frac{\operatorname{pr}\left(m_{i} \mid x\right)}{\operatorname{pr}\left(m_{j} \mid x\right)}=\frac{\frac{\operatorname{pr}\left(x \mid m_{i}\right) \operatorname{pr}\left(m_{i}\right)}{\operatorname{pr}(x)}}{\frac{\operatorname{pr}\left(x \mid m_{j}\right) \operatorname{pr}\left(m_{j}\right)}{\operatorname{pr}(x)}}=\frac{\operatorname{pr}\left(x \mid m_{i}\right) \operatorname{pr}\left(m_{i}\right)}{\operatorname{pr}\left(x \mid m_{j}\right) \operatorname{pr}\left(m_{j}\right)}
$$

because $\operatorname{pr}(x)$ drops

## MODEL SELECTION APPROACHES - BAYES FACTOR

- With a uniform prior over models we arrive at the Bayes factor [Kass and Raftery, 1995]

$$
\frac{\operatorname{pr}\left(x \mid m_{i}\right)}{\operatorname{pr}\left(x \mid m_{j}\right)}
$$

■ Hence, in Bayesian model selection, we evaluate a model $m$ based on its marginal likelihood

$$
\operatorname{pr}(x \mid m)=\int_{\theta} \operatorname{pr}(x \mid \theta, m) \operatorname{pr}(\theta \mid m) \mathrm{d} \theta
$$

where $\theta$ are the model parameters
■ The marginal likelihood is often difficult to evaluate, even numerically!

## Model selection approaches - BIC

- The marginal likelihood is tractable only for very simple models
- As an alternative, we use approximations of the marginal likelihood
- The Bayes information criterion (BIC) is such an approximation. Let $x$ contain $n$ samples and assume that $n \gg p$, then

$$
\begin{aligned}
\operatorname{pr}(x \mid m) & \approx \exp \left\{-\frac{1}{2} \operatorname{BIC}(x ; m)\right\} \\
\operatorname{BIC}(x ; m) & =-2 \log \operatorname{pr}(x \mid \hat{\theta}, m)+p \log (n)
\end{aligned}
$$

where $\hat{\theta}$ refers to the maximum likeklihood estimate and $p$ to the number of parameters

## MODEL SELECTION APPROACHES - BIC

- Let $\mathbf{Y}$ and $\epsilon$ be two random variables such that $\mathbf{Y}=f(X)+\epsilon$

■ Let $f_{\hat{\theta}}$ denote a maximum likelihood estimate on some training data

■ For $\epsilon \sim \operatorname{Normal}\left(0, \sigma^{2}\right)$ the BIC is related to the mean squared error with complexity penalty

$$
\begin{aligned}
\operatorname{BIC}(x ; m) & =\frac{1}{\sigma^{2}} \sum_{i=1}^{n}\left(y_{i}-f_{\hat{\theta}}\left(x_{i}\right)\right)^{2}+p \log (n)+C_{n} \\
& \propto \frac{1}{\sigma^{2}}\left\|y-f_{\hat{\theta}}(x)\right\|_{2}^{2}+p \log (n)
\end{aligned}
$$

where $C_{n}$ is a constant depending on $n$, which can be dropped for model comparison

## MODEL SELECTION APPROACHES - FIC

- BIC assumes $n \gg p$ and therefore depends only on the number of parameters

■ Fisher Information Approximation (FIA) [Ly et al., 2017]:

$$
\begin{aligned}
\operatorname{pr}(x \mid m) & \approx \exp \{-\operatorname{FIA}(x ; m)\} \\
\operatorname{FIA}(x ; m) & =\underbrace{-\log \operatorname{pr}(x \mid \hat{\theta}, m)+\frac{p}{2} \log \left(\frac{n}{2 \pi}\right)}_{\text {BIC like term }}+\log C_{m} \\
C_{m} & =\underbrace{\int_{\theta} \sqrt{\operatorname{det} \mathcal{I}_{m}(\theta)} \mathrm{d} \theta}_{\text {Geometric complexity }}
\end{aligned}
$$

where $\mathcal{I}_{m}$ denotes the Fisher information matrix

- $C_{m}$ is essential if $n \gg p$ is not given
[Cheema and Sugiyama, 2020]


## How do we control model complexity?

■ Regularization (e.g. ridge regression):

- Constrain the feasible set of parameter values
- Keep the number of parameters in the model constant, but allow them to become zero

■ Number of parameters:

- A good approximation of model complexity if $n<p$
- For $n>p$ we saw that the optimization problem has many solutions
- In deep neural networks, the gradient descent method can act similar to a regularizer
- Model complexity can decrease when adding more parameters (double descent)


## REGULARIZATION

## $l_{k}$-Penalized Regression

Objective function

$$
\begin{array}{ll}
\omega(\theta)=-\log \operatorname{pr}_{\theta}(y) & \text { (maximum likelihood), or } \\
\omega(\theta)=\|y-X \theta\|_{2}^{2} & \text { (linear regression) }
\end{array}
$$

Regularized estimate with $\ell_{k}$-norm penalty

$$
\hat{\theta}= \begin{cases}\underset{\theta}{\arg \min } & \omega(\theta) \\ \text { subject to } & \|\theta\|_{k}^{k}=\Lambda\end{cases}
$$

where

$$
\|\theta\|_{k}=\left(\sum_{j=2}^{p}\left|\theta_{j}\right|^{k}\right)^{1 / k}
$$

${ }^{2}$ Remember that we do not regularize the bias or $y$-intercept $\theta_{0}$

## $l_{k}$-PENALIZED REGRESSION

Identify saddle points of Lagrangian

$$
\mathcal{L}(\theta, \lambda)=\omega(\theta)+\lambda\left(\|\theta\|_{k}^{k}-\Lambda\right)
$$

In practice, we do not work with $\Lambda$, but set $\lambda$ such that the classification performance is optimal, i.e. we work with the Lagriangian

$$
\hat{\theta}(\lambda)=\underset{\theta}{\arg \min } \omega(\theta)+\lambda\|\theta\|_{k}^{k}
$$

At the optimum we must have

$$
\nabla_{\theta} \omega(\theta)+\lambda \nabla_{\theta}\|\theta\|_{k}^{k}=0
$$

i.e. the gradients of $\omega(\theta)$ and $\lambda\|\theta\|_{k}^{k}$ must point to opposite directions





## IMPLICIT REGULARIZATION AND DOUBLE DESCENT



Number of parameters


## MINIMUM $\ell_{2}$-NORM ESTIMATE - DF



${ }^{2}$ Requires a more advanced definition of DF that treats $X$ as random variable [Luan et al., 2021]

## IMPLICIT REGULARIZATION





Figure: Fitting degree $d=p-1$ Legendre polynomials. For $p>n$ the solution with the smallest $\ell_{2}$-norm is used.

[^1]
## TAKE HOME MESSAGES

■ Expected performance is the sum of training performance and model complexity

■ Complex models require regularization to prevent overfitting
■ The number of parameters does not correspont to the complexity of a model

■ Increasing the number of features can reduce model complexity if a min- $\ell_{2}$-norm estimator is used

■ If we have complex data and cannot make any assumptions on the generating process, we might be better off with an overparametrized model using regularization (success behind deep learning)

## More references

■ Akaike information criterion (AIC) [Akaike, 1974, Cavanaugh and Neath, 2019]

■ Bayesian information criterion (BIC) [Schwarz, 1978]

- Deviance information criterion (DIC) [Spiegelhalter et al., 2002]

■ Fisher Information Approximation (FIA) [Rissanen, 1996, Grünwald, 2007, Cheema and Sugiyama, 2020]

- Degrees of freedom (DF) [Tibshirani, 2015, Gao and Jojic, 2016, Luan et al., 2021]

■ Implicit regularization and double descent [Hastie et al., 2022, Luan et al., 2021, Derezinski et al., 2020, Kobak et al., 2020]

■ Sections 3.4, 7.3, 7.6, 7.7 and 7.9 [Hastie et al., 2009]

## The End

"All models are wrong, but some are useful." [Moody, 1991]

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[^0]:    ${ }^{1}$ The no free lunch theorem [Wolpert and Macready, 1997] tells us that there exists no generic model that works well on all domains, but we need to tailor our models to the data at hand in order to introduce a model bias, which reduces variance.

[^1]:    ${ }^{2}$ Legendre polynomials are quite useful, since their absolute value is bounded by one.

